

Maximum likelihood of Gaussian.

$$X \sim N(\mu, \sigma)$$

$$P(X|\theta) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left[-\frac{1}{2\sigma^2}(X-\mu)^2\right]$$

for $x_1, \dots, x_N$

$$\begin{aligned} P(x_1, \dots, x_N | \theta) &= \prod_{i=1}^N \frac{1}{(\sqrt{2\pi}\sigma)^2} \exp\left[-\frac{1}{2\sigma^2}(x_i - \mu)^2\right] \\ &= \frac{1}{(2\pi\sigma^2)^{N/2}} \exp\left(-\sum_{i=1}^N \frac{1}{2\sigma^2}(x_i - \mu)^2\right) \end{aligned}$$

Given log likelihood.

$$\begin{aligned} l(\theta, x) &= \log(P(x|\theta)) \\ &= \log\left[\frac{1}{(2\pi\sigma^2)^{N/2}} \exp\left(-\sum_{i=1}^N \frac{1}{2\sigma^2}(x_i - \mu)^2\right)\right] \\ &= -\frac{N}{2}(\log(2\pi) + \log\sigma^2) - \sum_{i=1}^N \frac{1}{2\sigma^2}(x_i - \mu)^2 \\ &= -\frac{N}{2} \log 2\pi - \frac{N}{2} \log \sigma^2 - \sum_{i=1}^N \frac{1}{2\sigma^2}(x_i - \mu)^2 \end{aligned}$$

$$\frac{\partial l}{\partial \mu} = 0 \Rightarrow \sum_{i=1}^N \frac{1}{\sigma^2}(x_i - \mu) = 0$$

$$\sum_{i=1}^N (x_i - \mu) = 0$$

$$\hat{\mu}_{ML} = \frac{\sum_{i=1}^N x_i}{N}$$

$$\begin{aligned} \frac{\partial l}{\partial \sigma^2} = 0 &\Rightarrow \left(\frac{1}{x}\right)' = -\frac{1}{x^2} \quad \begin{cases} \ln x = \frac{x}{1} \\ \log x = \frac{1}{x} \cdot \log e \end{cases} \\ &= -\frac{N}{2\sigma^2} + \frac{1}{2} \cdot \frac{1}{\sigma^4} \sum_{i=1}^N (x_i - \mu)^2 = 0 \end{aligned}$$

$$\Rightarrow \frac{1}{2\sigma^4} \sum_{i=1}^N (x_i - \mu)^2 = \frac{N}{2\sigma^2}$$

$$\frac{1}{\sigma^2} \sum_{i=1}^N (x_i - \mu)^2 = N$$

$$\sigma^2 = \frac{1}{N} \sum_{i=1}^N (x_i - \mu)^2$$

$$\sigma^2 = \frac{1}{N} \sum_{i=1}^N (x_i - \mu_{ML})^2$$

$$N(x|\mu, \Sigma) = \frac{1}{(2\pi)^{D/2} |\Sigma|^{1/2}} \exp\left(-\frac{1}{2} (x-\mu)^T \Sigma^{-1} (x-\mu)\right)$$

multivariate Gaussian.

$$\theta = (\theta_1, \dots, \theta_n)$$

$$X\theta = y$$

$$x_1 (x_{11} \dots x_{1n}) \rightarrow y_1$$

$$x_2 (x_{21} \dots x_{2n}) \rightarrow y_2$$

$$x_N (x_{N1} \dots x_{Nn}) \rightarrow y_N$$

$$\begin{matrix} \xrightarrow{n} \\ \downarrow N \end{matrix} \begin{pmatrix} x_1 & \dots & x_n \\ x_2 & \dots & x_n \\ \vdots & & \vdots \\ x_N & \dots & x_n \end{pmatrix} \begin{pmatrix} \theta_1 \\ \vdots \\ \theta_n \end{pmatrix} = \begin{pmatrix} y_1 \\ y_2 \\ \vdots \\ y_N \end{pmatrix}$$

$$E(\theta) = (y - A\theta)^T (y - A\theta)$$

$$= (y^T - \theta^T A^T) (y - A\theta)$$

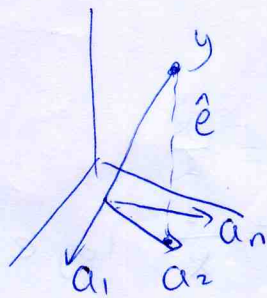
$$= y^T y + \theta^T A^T A \theta - y^T A \theta - \theta^T A^T y$$

$$= y^T y + \theta^T A^T A \theta - \underbrace{2 y^T A \theta}_{-2 \theta^T A^T y} \text{ (Scalar)}$$

$$\frac{\partial E}{\partial \theta} = 2 A^T A \theta - 2 y^T A = 0$$

$$A^T A \theta = y^T A \Rightarrow \hat{\theta} = (A^T A)^{-1} y^T A$$

$$\frac{\partial E}{\partial \theta} = 0 \Rightarrow \hat{\theta} = (A^T A)^{-1} A^T y$$



$$\vec{a}_1 \theta_1 + \vec{a}_2 \theta_2 + \dots + \vec{a}_n \theta_n = y$$

$$A^T (y - A\hat{\theta}) = 0$$

$$A^T y - A^T A \hat{\theta} = 0$$

$$\hat{\theta} = (A^T A)^{-1} A^T y$$

$$\begin{pmatrix} y_1 \\ \vdots \\ y_N \end{pmatrix} = \begin{pmatrix} A\theta_1 \\ \vdots \\ A\theta_N \end{pmatrix}$$

Stochastic process

A collection of random variable. $X_1, \dots, X_n$

Chapman - Kolmogorov Equation.

Suppose that $\{f_i\}$ is an indexed collection of random variables. that is, a ~~st~~ stochastic process

Let $P_{i_1, \dots, i_n}(f_1, \dots, f_n)$ be the joint probability density function of the values of the random variables $f_1, \dots, f_n$. Then, the Chapman - Kolmogorov equation is:

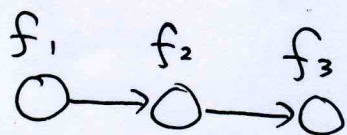
$$P_{i_1, \dots, i_{n-1}}(f_1, \dots, f_{n-1}) = \int_{-\infty}^{\infty} P_{i_1, \dots, i_n}(f_1, \dots, f_n) df_n$$

Given a Markov chain:

$$P_{i_1, \dots, i_n}(f_1, \dots, f_n) = P_{i_1}(f_1) P_{i_2, i_1}(f_2 | f_1) \dots$$

$P_{i_j, i_{j-1}}(f_j | f_{j-1})$ is called the transition probability.

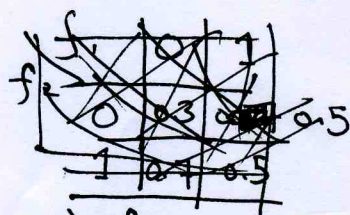
$$P_{i_3, i_1}(f_3 | f_1) = \int_{-\infty}^{\infty} P_{i_3, i_2}(f_3 | f_2) P_{i_2, i_1}(f_2 | f_1) df_2$$



$$P(f_3=1 | f_1=0) = ?$$

$$\begin{aligned} &P(f_3=1 | f_2=0) P(f_2=0 | f_1=0) + \\ &P(f_3=1 | f_2=1) P(f_2=1 | f_1=0) \\ &= 0.3 \times 0.9 + 0.7 \times 0.2 \\ &= 0.27 + 0.14 = 0.41 \end{aligned}$$

$$\Rightarrow P_{09} =$$



$f_1 \backslash f_2$	0	1
0	0.3	0.7
1	0.4	0.6

$f_2 \backslash f_3$	0	1
0	0.1	0.9
1	0.8	0.2

Example of Markov chain.

$$P(X_n | X_{n-1} \dots X_1) = P(X_n | X_{n-1})$$

$$P_{ij}(t) \stackrel{\Delta}{=} P(X_{t+1}=j | X_t=i)$$

is the transition probability from state i to state j at time t . If the transition probability does not depend on time.

$$P_{ij}(t) \rightarrow P_{ij}$$

We say that the Markov chain is time-homogeneous.

From DNA to transitional matrix
2-gram language model.



{ ACATCCTTTGGCGTC

4x4

, I am a student of Chinese university |V| x |V|

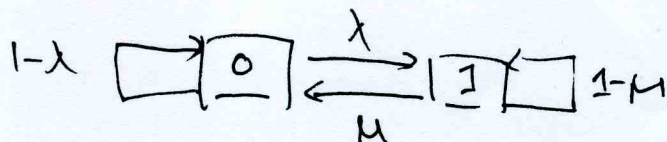
PATH probability :

the conditional probability of a path, conditional on the first value, is the product of the transitional probabilities between successive states of the path.

$$P(X_1, X_2 \dots X_t) = (X_1, X_2 \dots X_t) = P(X_1=x_1) P_{x_1 x_2} P_{x_2 x_3} \dots P_{x_{t-1} x_t}$$

$$P(X_1=x_1, X_2=x_2, X_3=x_3) = P(X_1=x_1) P_{x_1 x_2} P_{x_2 x_3} = P_{x_1} P_{x_1 x_2} P_{x_2 x_3}$$

$$P(X_3=x_3 | X_1=x_1) = \sum_{x_2} P_{x_1 x_2} P_{x_2 x_3}$$



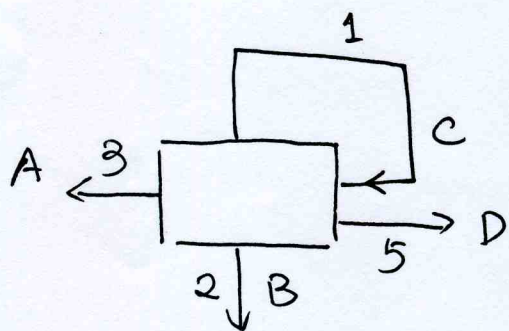
$$A = \begin{pmatrix} P_{00} & P_{01} \\ P_{10} & P_{11} \end{pmatrix} = \begin{pmatrix} 1-\lambda & \lambda \\ \mu & 1-\mu \end{pmatrix}$$

$$P(X_1=0) = 0.5$$

$$P(0, 0, 0, 1, 1, 0) = ?$$

$$= 0.5 P_{00} P_{00} P_{01} P_{11} P_{10} = 0.5 (1-\lambda) (1-\lambda) \lambda (1-\mu) \mu$$

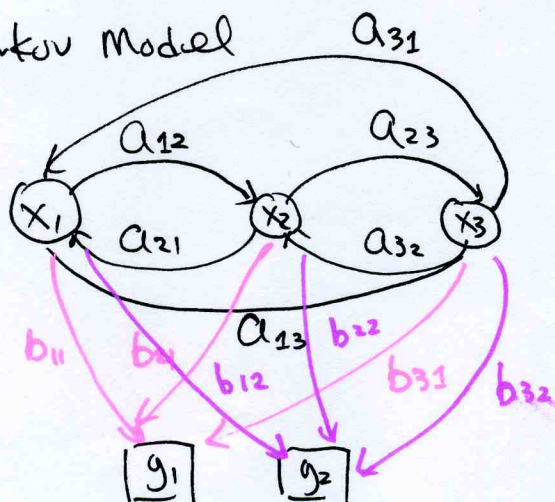
Similarly, we can calculate the probability of a sequence of DNA sequence GGCTTTCAAGCT ... with $(\frac{1}{4}, \frac{1}{4}, \frac{1}{4}, \frac{1}{4})$



$$Ave = \frac{1}{4} \times 3 + \frac{1}{4} \times 2 + \frac{1}{4} \times 5 + \frac{1}{4} \times (Ave + 1)$$

$$\Rightarrow Ave = ??$$

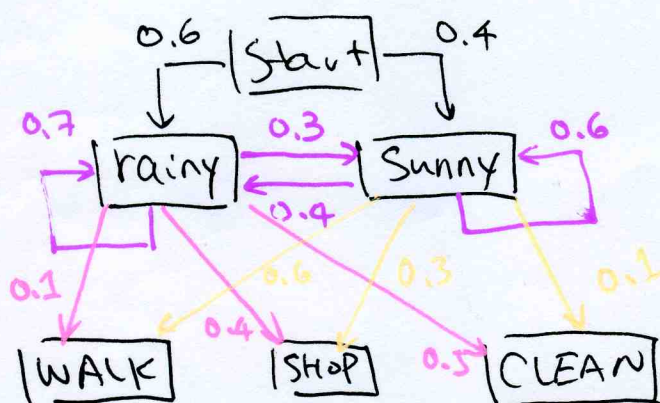
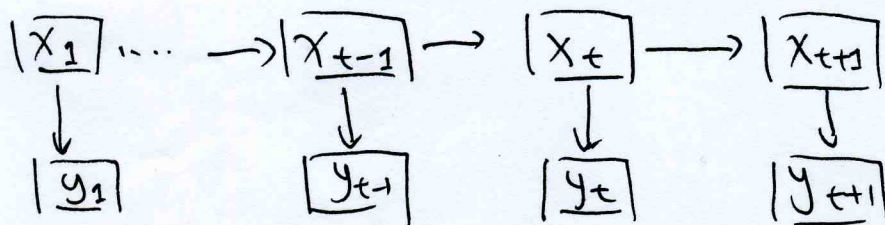
Hidden Markov Model



$$A = \begin{pmatrix} a_{11} & a_{12} & a_{13} \\ a_{21} & a_{22} & a_{23} \\ a_{31} & a_{32} & a_{33} \end{pmatrix}$$

$$B = \begin{pmatrix} b_{11} & b_{12} \\ b_{21} & b_{22} \\ b_{31} & b_{32} \end{pmatrix}$$

Given $(y_t, y_{t-1}, \dots, y_1)$ how to estimate A and B.



Most likely sequence can be

Viterbi Algorithm

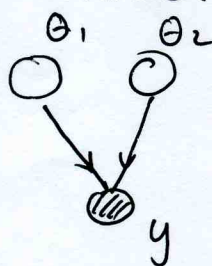
Given a HMM with state X , initial probability π_i of being in state i and transition probabilities a_{ij} from state i to state j . Observe the outputs $y_0, \dots, y_T$. What is the most likely sequence of states $(x_0, \dots, x_T)$?

$$\begin{cases} V_{0k} = P(y_0 | k) \pi_k \\ V_{tk} = P(y_t | k) \cdot \max_{x \in X} (a_{x,k} V_{t-1,x}) \end{cases} \quad \begin{cases} \pi_0 = \{0.6, 0.4\} \\ X = \{\text{rainy}, \text{sunny}\} \end{cases}$$

V_{tk} is the probability of the most probable state sequence responsible the first $t+1$ observations. that has the k as the final state. Let $P_{tr}(k, t)$ be the function that returns the value of x used to compute $V_{t,k}$ if $t > 0$, or k if $t = 0$. Then:

$$\begin{cases} X_T = \operatorname{argmax}_{x \in X} (V_T, x) \\ X_{t-1} = P_{tr}(X_t, t) \end{cases}$$

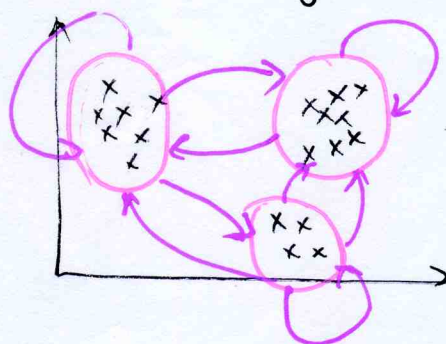
The similarity to mixture model. HMM is a dynamic mixture model.



θ_1 - State 1

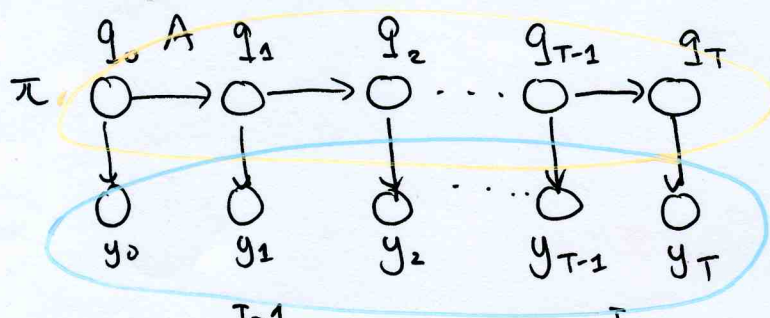
θ_2 - State 2.

{ emission probability is the direct probability distribution.



It has the state transition probability. And the data is generated in sequence.

New Notation of Jordan



q - state

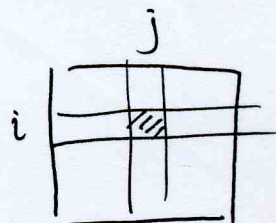
y - output

A - homogeneous transition matrix.

$$P(q, y) = P(q_0) \prod_{t=0}^{T-1} P(q_{t+1} | q_t) \prod_{t=0}^{T-1} P(y_t | q_t)$$

$$a_{q_t, q_{t+1}} = \prod_{i,j=1}^M [a_{ij}]^{q_t^i q_{t+1}^j} \Rightarrow$$

$$P(q, y) = \pi_{q_0} \prod_{t=0}^{T-1} a_{q_t q_{t+1}} \prod_{t=0}^{T-1} P(y_t | q_t)$$



$$P(y) = \sum_q P(q, y) = \sum_{q_0} \sum_{q_1} \dots \sum_{q_T} \pi_{q_0} \prod_{t=0}^T A_{q_t q_{t+1}} \prod_{t=0}^T P(y_{t+1} | q_t)$$

$$P(q_t | y) = \frac{P(y | q_t) P(q_t)}{P(y)}$$

$$= \frac{P(y_0 \dots y_t | q_t) P(q_t) P(y_{t+1} \dots y_T | q_t)}{P(y)}$$

$$= \frac{\alpha(q_t) \beta(q_t)}{P(y)}$$

$$\alpha(q_t) \triangleq P(y_0 \dots y_t, q_t)$$

$$\beta(q_t) \triangleq P(y_{t+1} \dots y_T | q_t)$$

$\alpha(q_t)$ is a joint probability while $\beta(q_t)$ is a conditional probability. 1, 2 and 3:

$$P(y) = \sum_{q_t} \alpha(q_t) \beta(q_t)$$

$$P(y) = \sum_{q_t} P(y, q_t)$$

$$r(q_t) \triangleq \frac{\alpha(q_t) \beta(q_t)}{P(y)}$$

A new notation of posterior probability.

$$\begin{aligned} \alpha(q_{t+1}) &= P(y_0, \dots, y_{t+1}, q_{t+1}) \\ &= P(y_0, \dots, y_{t+1} | q_{t+1}) P(q_{t+1}) \\ &= P(y_0, \dots, y_t | q_{t+1}) P(y_{t+1} | q_{t+1}) P(q_{t+1}) \\ &= P(y_0, \dots, y_t, q_{t+1}) P(y_{t+1} | q_{t+1}) \\ &= \sum_{q_t} P(y_0, \dots, y_t, q_t, q_{t+1}) P(y_{t+1} | q_{t+1}) \\ &= \sum_{q_t} P(y_0, \dots, y_t, q_{t+1} | q_t) P(q_t) P(y_{t+1} | q_{t+1}) \\ &= \sum_{q_t} P(y_0, \dots, y_t | q_t) P(q_{t+1} | q_t) P(q_t) P(y_{t+1} | q_{t+1}) \\ &= \sum_{q_t} P(y_0, \dots, y_t, q_t) P(q_{t+1} | q_t) P(y_{t+1} | q_{t+1}) \\ &= \sum_{q_t} \alpha(q_t) A_{q_t q_{t+1}} P(y_{t+1} | q_{t+1}) \end{aligned}$$

$$\alpha(q_0) = P(y_0, q_0) = P(y_0 | q_0) P(q_0) = P(y_0 | q_0) \pi_{q_0}$$

$$\begin{aligned} \beta(q_t) &= P(y_{t+1} \dots y_T | q_t) = \sum_{q_{t+1}} P(y_{t+1} \dots y_T, q_{t+1} | q_t) \\ &= \sum_{q_{t+1}} P(y_{t+1} | q_{t+1}) P(y_{t+2} \dots y_T | q_{t+1}) P(q_{t+1} | q_t) \\ &= \sum_{q_{t+1}} \beta(q_{t+1}) A_{q_t q_{t+1}} P(y_{t+1} | q_{t+1}) \end{aligned}$$